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1	Chemometric-Assisted Cocrystallization: Supervised Pattern Recognition for Predicting the
2	Formation of New Functional Cocrystals
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27 Abstract

Owing to the antimicrobial and insecticide properties, the use of natural compounds like essential oils and their active components has proven to be an effective alternative to synthetic chemicals in different fields ranging- from drug delivery to agriculture and from nutrition to food preservation. Their limited application due to the high volatility and scarce water solubility can be expanded by using crystal engineering approaches to tune some properties of the active molecule by combining it with a suitable partner molecule (coformer). However, the selection of coformers and the experimental effort required for discovering cocrystals are the bottleneck of cocrystal engineering. This study explores the use of chemometrics to aid the discovery of cocrystals of active ingredients suitable for various applications. Partial Least Squares-Discriminant Analysis is used to discern cocrystals from binary mixtures based on the molecular features of the coformers. For the first time, by including failed cocrystallization data and considering a variety of chemically diverse compounds, the proposed method resulted in a successful prediction rate of $\frac{8385}{5}\%$ for the test set in the model validation phase and of $\frac{6274}{5}\%$ for the external test set.

50 Keywords

51 cocrystal, crystal engineering, chemoinformatics, chemometrics, partial least square discriminant analysis, Quantitative

52 Structure–Property Relationship

53 List of abbre	eviations (sorted	alphabetically)
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- 54 <u>ACC_%; Accuracy</u>
- 55 BM: Binary Mixture
- 56 CC: Cocrystal
- 57 CCDC: Cambridge Crystallographic Data Centre
- 58 CSD: Cambridge Structural Database
- 59 EO: Essential Oil
- 60 FDA: Food and Drug Administration
- 61 GRAS: Generally Recognized As Safe
- 62 LV: Latent Variable
- 63 MEP: Molecular Electrostatic Potential
- 64 NER_%: Non Error classification Rate
- 65 PC: Principal Component
- 66 PCA: Principal Component Analysis
- 67 PLS–DA: Partial Least Squares–Discriminant Analysis
- 68 PXRD: Powder X-Ray Diffraction
- 69 QSPR: Quantitative Structure–Property Relationship
- 70 <u>SEN_%: Sensitivity</u>
- 71 VIP: Variable Importance in Projection.
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82 **1. Introduction**

In the last few decades, the use of agrochemicals and food preservatives has grown exponentially as a direct consequence of the rapid increase of the world population [1,2]. Owing to their potential adverse effects on both human health and environment [2–4], alternative strategies based on the use of more sustainable chemicals have been proposed to support the food system. Being able to exert antimicrobial, insecticidal, and antioxidant properties [5,6], essential oils (EOs) and their active components have been used as green substitutes of synthetic chemicals to extend the shelf-life of foodstuff and in pests control [7–9]. These compounds are Generally Recognized As Safe (GRAS) by the Food and Drug Administration (FDA) [10], however, despite their appealing properties, their use is limited by their high volatility and poor stability [7,9,11,12].

In fact, physicochemical properties of materials play a key role in determining whether a chemical is suitable for a specific purpose, thus strongly affecting its field of application. Scientists have always desired to obtain materials with target properties, and crystal engineering is one of the most interesting approaches to synthesize a great variety of crystalline materials for applications in various fields, ranging from pharmaceuticals to agrochemicals, and from nutraceuticals to cosmetics [13–16]. The basic idea of crystal engineering is related to the possibility of controlling the crystal structure of molecules and, therefore, the properties of the resulting solids. Polymorphism, vitrification and cocrystallization are some of the available strategies to modify the intrinsic properties of molecules without the need of synthetic modifications [17–21].

97 Cocrystals are multicomponent crystalline solid materials in which the constituents (i.e., coformers) are bound in a well-98 defined stoichiometric ratio [22,23] *via* non-covalent interactions (e.g., hydrogen bonds, halogen bonds, π - π stacking) within 99 the same crystal structure. Cocrystallization allows for the combination of the desired molecule of interest with properly 100 selected partner molecules, paving the way to an array of potential materials with enhanced properties [19,24,25]. Within this 101 frame of reference, cocrystals based on the active components of EOs have been proposed as active ingredients for food 102 packaging, agrochemical and pharmaceutical applications [7,19,26,27].

Despite the great advantages offered by cocrystallization, the proper degree of complementarity between the two partner molecules required to obtain crystalline materials with the desired properties is not easy to assess [28–31]. In this context, the selection of coformers and the great effort required for both the systematic experimental screening and careful characterization of the products derived from the combination of all the possible coformer pairs represent the major bottleneck of cocrystal engineering. Computational techniques represent a powerful tool to reduce the experimental effort required for the discovery of new cocrystals, enabling to evaluate beforehand whether a cocrystal can be obtained starting from pre-selected coformers. 109 These *in silico* strategies can be based on the calculation of a variety of parameters useful to predict the formation of a 110 cocrystal, such as lattice energy [32], solubility [33], hydrogen bond propensity along with the quantitation of molecular 111 interaction energy [29,30], and molecular complementarity [34].

Despite the massive efforts spent to develop a method to predict cocrystal formation, at present none of the proposed strategieshas proven to be both totally reliable and easy to apply.

114 Chemometrics could play a pivotal role in cocrystal discovery: up to now only few Machine Learning methods have been 115 proposed in predicting cocrystal formation, enabling the screening of new cocrystals once a supervised model is properly 116 trained and validated. In the study proposed by Devogelaer et al., information of successful cocrystallization experiments was directly taken from the Cambridge Structural Database (CSD) [35] and Artificial Neural Networks (ANN) were used to predict 117 118 the formation of new cocrystals [36]. Similarly, Wang et al. relied on a consensus method based on multiple Random Forest 119 algorithms, in which the successful cocrystallization dataset was integrated with randomly generated failed cocrystallization 120 data [37]. These approaches are reported in the literature as network-based methods. Additional studies were based on the use 121 of successful and unsuccessful cocrystallization datasets obtained from experimentation, literature, and/or the CSD for 122 screening specific classes of coformers. Within this framework, Przybyłek et al. used Multivariate Adaptive Regression 123 Splines to predict the formation of dicarboxylic and phenolic acid-based cocrystals [38,39], whereas Wicker et al. focused on 124 variously substituted benzoic acids and benzamides using a Support Vector Machine algorithm [24]. Vriza et al. used an 125 ensemble one-class classification method to aid the discovery of $\pi - \pi$ corrystals, thus giving a great contribution in enriching 126 one of the most under-represented classes of cocrystals in the CSD [40]. Most recently, Mswahili et al. developed a cocrystal 127 screening method based on ANN by using both successful and unsuccessful experimental cocrystallization data retrieved from 128 the literature and a plethora of molecular descriptors calculated using Mordred [41,42].

In the frame of a research activity dealing with the synthesis of new functional cocrystals based on the active constituents of EOs and other GRAS molecules to broaden their applicability in the industrial field [7,19], we propose a chemometric approach to aid the discovery of new cocrystalline materials.

For the first time, a training set based on the results of failed (binary mixtures, BM) and successful (cocrystal, CC) cocrystallization experiments was used for the computation of a Quantitative Structure–Property Relationship-like (QSPR) model based on Partial Least Squares–Discriminant Analysis (PLS–DA), after preliminary exploratory analysis by Principal Component Analysis (PCA). The PLS-DA approach, with respect to network-based methods offers the advantages of having

- 136 only one parameter to optimize, i.e., the model dimensionality, and the direct interpretation of the importance of descriptors
- 137 in classification, while highlighting their interplay (by inspection of weights and loadings plots).
- 138 The effectiveness of the study relies on the use of compounds belonging to different chemical classes and a reduced number
- 139 of 1D, 2D, and 3D molecular descriptors of various nature (e.g., constitutional, geometric, physical, topological, and surface
- area-based descriptors) [24,38,39,43], enabling the high-throughput screening of novel cocrystalline materials and offering
- 141 maximum flexibility and effectiveness at a minimum computational and experimental cost.
- 142

143 2. Experimental Procedures

144 2.1. Mechanochemical protocol and class assignation

All the molecules in the dataset were chosen among the list of GRAS molecules drawn up by the FDA [10]. Selected pair of molecules among the chosen ones were assigned either to the CC class or to the BM class. Pairs of molecules in the dataset for which a cocrystal structure was already described in literature were individuated in the Cambridge Structural Database (CSD) [35] with the Cambridge Crystallographic Data Centre (CCDC) software ConQuest [44] and visualized with Mercury [45]. They are reported in Section 3 of the Supplementary Material, together with their unique CSD refcode and the reference to the original publications.

151 Cocrystallization for all the pairs with no known structure in literature was instead attempted with the following 152 mechanochemical protocol. All the reagents employed were commercially available and used as such in all the experiments. 153 Equimolar amounts of each reagent were directly mixed in an agate mortar and subjected to manual grinding for 10-15 154 minutes, without using any solvent. The resulting powder samples were collected in closed vials. Assignation to CC or BM 155 classes was performed by comparing the Powder X-ray Diffraction (PXRD) pattern of the ground sample with those of the 156 pure reagents. Possible occurrence of polymorphic transitions for the reagents was excluded by comparing the experimental 157 PXRD data after milling with the calculated pattern of all the known crystalline forms of the reagents. The occurrence of new 158 peaks, unexplained by the presence of unreacted reagents, was taken as indication that cocrystallization had occurred and the 159 sample was assigned to the CC class. In case no additional peaks appeared in the PXRD pattern the sample was instead 160 classified as a BM.

- 161
- 162

163 2.2. Powder X-ray diffraction

164 Typically, PXRD data were collected on a Rigaku Smartlab XE diffractometer in θ - θ Bragg-Brentano geometry with Cu K α 165 radiation. The samples were placed on glass supports and exposed to radiation $(1.5^{\circ} \le 2\theta \le 50^{\circ})$ at a scan rate of 10°/min. The 166 diffracted beam was collected on a 2D Hypix 3000 solid state detector. 5° radiant soller were used as a compromise for high 167 flux and moderate peak asymmetry at low angles. Beam stopper and anti-scatterer air component were used to mitigate the 168 profile at low angle. In some rare cases, the data were collected on a Thermo Fisher Scientific ARL X'TRA diffractometer in 169 θ - θ Bragg-Brentano geometry with Cu K α radiation (3° $\le 2\theta \le 30^{\circ}$ at a scan rate of 5°/min, or 3° $\le 2\theta \le 40^{\circ}$ at a scan rate of 170 0.3°/min).

171

3. Computational Methods

173 3.1. Molecular descriptors calculation

For each molecule 31 molecular descriptors were calculated (Table S1). A theoretical background for the less knowndescriptors is given in Section 4 of the Supplementary Material.

176 The molecular weight, the number of atoms, the number of bonds, the number of hydrogen bond donor sites and the number 177 of hydrogen bond acceptor sites were calculated with FLAP software (Fingerprint for Ligand and Protein) [46] at pH 7.0, 178 using the 3D structures of all molecules in SDF format as input (downloaded from the PubChem database). The number of 179 rotatable bonds, the number of rings, the hydrophobicity (accounted as the number of hydrophobic centers), the logP 180 (logarithm of octanol/water partition coefficient), the molecular volume, the total molecular dipole moment (based on point 181 charge distribution in the molecule), and its components along the axes (using the principal axes of the molecular graph) were 182 then calculated for the same structures using Sybyl 8.1 [47] (www.tripos.com) and taking in consideration the protonation 183 state of molecules. The same software was also used to estimate the strain energy of the molecule without performing any 184 geometry optimization. This energy term relies on an electrostatic calculation from atomic charges using the internal Tripos 185 force field [48]. For the estimation of molecular volume and dipole moment, a specific SPL script was employed. The 186 calculated volume is enclosed in a water-accessible surface computed at a repulsive interaction energy of 0.20 kcal/mol with 187 a water probe. A custom Python script was used to automatically calculate the Solvent Accessible Surface Area (SASA) in 188 PyMol 2.0 [49], with the dot density parameter set to 4. The number of heteroatoms, the number of valence electrons, and the 189 indexes ${}^{0}\chi$, ${}^{0}\chi^{\nu}$,

190 the open-source cheminformatics toolkit RDKit Q4 2013 [50]. The average isotropic polarizability α_{iso} , the polarizability 191 anisotropy $\Delta \alpha$, and the Molecular Electrostatic Potential (MEP) were calculated with Gaussian 16 [51] following the *in vacuo* 192 Density-Functional Theory optimization of all the molecules, employing the hybrid functional B3LYP and the People double-193 z basis set 6-31+g(d,p).

194 Postprocessing of the MEP to extract critical points at a given electron density isosurface was performed with a custom Python 195 3.6.1 script on a three-dimensional map (cube format) with a sampling density of 6 points/Bohr along the three directions. 196 The MEP was analyzed at an electron density isosurface of 0.002 a.u. with a tolerance of 0.001 a.u., meaning that only MEP 197 values corresponding to regions of space with electron density in the 0.001–0.003 a.u. range were considered. A first set of critical points was identified comparing MEP values of each cube point with those of its 6 nearest neighbors. A point was 198 199 considered a local minimum if the number of nearest neighbors with higher MEP was greater or equal to a given integer (4). 200 Likewise, a point was considered a local maximum if the number of nearest neighbors with lower MEP was greater or equal 201 to the same integer. This first step yielded a large number of candidate critical points encompassing a wide range of MEP 202 values. Since our focus was on identifying the regions of the molecules likely to be involved in strong hydrogen bonds within 203 the cocrystal, in a second step this first set of points was filtered based on the MEP values of the global minimum and 204 maximum. This was done as follows: for each local minimum (maximum), the ratio between its MEP value and that of the 205 global minimum (maximum) was computed, and the point was kept only if the ratio exceeded a given threshold (0.1). In this 206 way, only points corresponding to shallow critical points were discarded. This step allowed to identify the MEP isosurface 207 regions corresponding to hydrogen bond donors and acceptors. However, due to the rugged character of the MEP map, 208 multiple critical points of the same type could appear in close proximity. To univocally map a given region of the isosurface 209 to a MEP value, critical points close to each other (below a distance threshold of 1.0 Bohr) were merged iteratively, keeping 210 only the lower MEP point for minima and higher MEP point for maxima. The algorithm then provided the final set of MEP 211 critical points at the given electron density isosurface.

212

213 3.2. Data analysis

The entire data analysis was carried out in MATLAB R2019a environment (Mathworks, Natick, Massachusetts, USA) with
 the aid of the- PLS_Toolbox 8.7.1 (Eigenvector Research Inc., Washington, USA) chemometric package.

3.2.1. Data preprocessing

homemade MATLAB routine.

Each sample was described by m = 31 variables (Table S1): the <u>absolute value of the difference between difference in absolute</u> value of the molecular descriptors of the two partner molecules was calculated, thus obtaining the predictor matrix X (181 × 31). The class membership was binary encoded (1: belonging to the class; 0: otherwise) in a dummy matrix Y (181 × 2) with each column representing one of the two modelled classes. The dataset was split in two subsets by using the <u>Kennard Stone</u> samplingDuplex algorithm [52]: <u>8070</u>% of the data were used as calibration set, X_{cal} (<u>146-127</u> × 31) and Y_{cal} (<u>146-127</u> × 2), whereas the remaining <u>2030</u>% were used as test set, X_{test} (<u>35-54</u> × 31) and Y_{test} (<u>(3554</u> × 2)). The calibration set and the test set are reported in Table S2 and Table S3, respectively.

-was used to carry out preprocessing, Principal Component Analysis (PCA) and PLS - DA computation, and to split the original

dataset into calibration and test set. The proper number of latent variables (LVs) to be retained was evaluated by running a

Before carrying out both exploratory multivariate data analysis and the computation of the supervised model, the calibration matrix X_{cal} was preprocessed column-wise by performing mean centering and scaling to unit variance. Mean centering was applied on the response matrix Y_{cal} to ensure the stability of the model.

232

233 *3.2.2. Exploratory multivariate data analysis*

PCA [53–55] was carried out preliminarily on the calibration set X_{cal} to assess the distribution of the samples and to check for potential data structures. Reduction of data dimensionality is carried out through the linear combination of the original variables in a set of orthogonal ones, i.e., Principal Components (PCs), which identify the direction of maximum variance. This is summarized in the decomposition equation:

238

$X_{\rm cal} = TP^{\rm T} + E$

where T and P represent, respectively, the coordinates of the samples projected in the reduced space, i.e., the scores, and the weights each original variable has on a given PC, i.e., the loadings. The deviations from the model are accounted in the error matrix E.

243 *3.2.3. Supervised pattern recognition*

PLS-DA [56,57] was used to discriminate pairs of partner molecules whose combination forms CCs from the ones giving BMs. PLS-DA is based on PLS regression [58]. Briefly, this supervised technique decomposes the predictor matrix X_{cal} and the dependent variables matrix Y_{cal} in a PCA-like way and imposes inner linear relationships between the X_{cal} and the Y_{cal} scores as follows:

248

U = bT

where T and U, are the X_{cal} and the Y_{cal} scores, respectively. This is accomplished by rotating the Latent Variable (LV) space of X_{cal} through a weight matrix W in a way that maximizes the covariance between T and U. The PLS regression model is summarized as:

where *E* is the error matrix and *B* is the pseudo-regression coefficient matrix expressed according to the following equation:

$$B = W(P^T W)^{-1} \operatorname{diag}(b) Q$$

where P and Q are the X_{cal} and the Y_{cal} loadings, respectively.

In this case, the dependent variables in the Y_{cal} matrix are defined as dummy variables, one for each modelled class, taking values of 1 if the sample belongs to the class and 0 otherwise. Current implementation of PLS–DA may differ on the basis of how the classification rule is defined. In this work, a pure discriminant rule (samples are assigned univocally to only one category) was applied, and thus a sample is assigned to the class for which the predicted response \hat{y} is the highest (i.e., \hat{Y} values are continuous and not dummy as they were codified).

- The proper number of LVs was chosen according to the maximum Non Error classification Rateaccuracy (NERACC_%; i.e., the percentage of samples correctly assigned to the respective class) in leave-more-out cross validation, adopting a Venetian blinds cancellation scheme with 10 splits (blind thickness: 1). This operation was carried out by running a custom MATLAB routine. The performance of the classification model was evaluated both on the calibration and the test sets in terms of NERACC_% as well as showing the confusion matrix. In addition, the sensitivity (SEN_%, i.e., the percentage of samples within a class that were correctly assigned to their class) was calculated for both classes.
- 267 The importance of each predictor was estimated in terms of Variable Importance in Projection (VIP score) [56]. The VIP
- score of the jth variable in the X space is defined as the component-wise sum of its PLS weight w_i on the fth component
- 269 multiplied by the fraction of variance of the *Y* explained by that component, according to the following equation:

270
$$VIP_j^2 = \frac{1}{SSY_{\text{tot}}F} \sum_{f=1}^F w_{jf}^2 SSY_f J$$

271 where J is the number of variables in the X space and F is the number of LVs that were retained. Since:

$$\sum_{j=1}^{J} VIP_j^2 = J$$

the proposed threshold for determining whether a variable could be considered important is set to 1..

274 Finally, the predictive capability of the model was evaluated on an external set of N = 58 binary combinations of partner

275 molecules. An overview of the involved samples is reported in Table S4 along with their estimated \hat{y} values.

276 <u>The number of entries in the BM and CC classes for training, test, and external validation sets is reported in Table 1.</u>

277

278 <u>Table 1. Number of CC and BM samples in the calibration, test, and external validation set. The last column reports the total</u>

279 <u>number N of samples per set.</u>

	<u>CC</u>	<u>BM</u>	<u>N</u>
Calibration set	<u>71</u>	<u>56</u>	<u>127</u>
<u>Test set</u>	<u>30</u>	<u>24</u>	<u>54</u>
External validation set	<u>31</u>	<u>27</u>	<u>58</u>

280

281 **4. Results and Discussion**

282 The cocrystallization experiments were carried out mechanochemically by manual neat grinding of the two substances. This

283 method was selected among many possible others due to its simplicity and promptness, allowing us to screen several molecular

pairs in a standardized way [21,59]. The classification as BM or CC was assessed by PXRD patterns (available in Section 3

of the Supplementary Material). Cocrystals already present in the CSD were also included in our dataset.

286

287 4.1. Data analysis

288 Data handling prior to analysis can affect the way the model is trained, thus having consequences on its interpretation.

- 289 Since samples are the result of the combination of two partner molecules, each one described by its own set of descriptors,
- 290 the concatenation strategy, i.e., listing the descriptors of the first coformer followed by those of the second coformer, was

discarded due to the lack of commutation between the two sets of descriptors. In fact, the order in which the molecular descriptors are listed represents an *a priori* decision on which compound is acting as molecule of interest or partner molecule. In our case this would be sub-optimal since many of the molecules in our dataset could assume both roles. Therefore, in order to address the problem described above a commutative strategy capable of avoiding the production of indeterminate forms should be chosen. Considering that in the dataset many constitutional molecular descriptors were characterized by a few nonzero values, the calculation of both products and ratios between the descriptors was discarded since additional zero values or indeterminate forms could be generated. Also, the division is non-commutative.

In order to combine the two partner molecules without imposition on their role, the <u>absolute value of the difference</u> difference in absolute value between the molecular descriptors of the partner molecules was calculated and used to describe each binary combination [38,39], giving maximum flexibility to the model. The information achieved is still relevant and easily interpretable since it is related to the dissimilarity between the descriptors. In fact, the differences in absolute value<u>absolute</u> value of the differences between descriptors for each case are the elements of the Manhattan distance [60], one of the possible indexes to account for the dissimilarity between cases in a multivariate way.

In the present study, the use of basic linear modelling methods was preferred with respect to non-linear modelling, such as
 ANN [61,62], to keep the calculations as simple as possible, and to ensure a certain degree of interpretability of the results.
 Furthermore, the number of samples is too limited to ensure proper tuning of the ANN hyperparameters.

For a fruitful discussion, samples and variables are reported in the text according to their identification number as follows: i)
samples are written in plain text; ii) variables are underlined. The key is available in Tables S1–S3.

309

310 *4.1.1. Exploratory multivariate data analysis*

311 After data preprocessing, PCA was used in an exploratory way to assess the presence of potential data structures in the 312 calibration set.

Four PCs were retained explaining 67% of the variance. As shown in the score plots (Figure 1), a mild segregation separation was present in the PC <u>3-2</u>*vs*. PC <u>4-3</u> score plot, with the groups separated by the bisecting line of the II and the <u>HI-IV</u> quadrant. In addition, most of the CC samples occupied the III quadrant of the PC 2 *vs*. PC 3 score plot and were, in general, less scattered than BM samples.



Figure 1. Scores of the samples on the first 4 PCs: PC 1 *vs.* PC 2 (left), PC 2 *vs.* PC 3 (center), PC 3 *vs.* PC 4 (right). The fraction of variance explained by a given component is reported as a percentage value in parenthesis on the corresponding axis. Samples are marked according to their class (empty-magenta diamonds: cocrystals; black-cyan circles: binary mixtures).

324 As for the loading plots depicted in Figure <u>S12</u>, it can be observed that PC 1 explains the features related both to differences 325 in molecular dimensions (e.g., <u>1</u>, <u>2</u>, <u>3</u>, <u>9</u>, <u>16</u>, <u>29</u>) and connectivity (<u>23</u>, <u>24</u>, <u>25</u>, <u>26</u>). PC 2 considers the dissimilarities in the 326 electronic properties (e.g., <u>6</u>, <u>7</u>, <u>19</u>, <u>22</u>, <u>30</u>) of the two partner molecules as well as the difference in their number of 327 heteroatoms (27). PC 3 accounts for more specific features, such as differences in molecular complexity (21), molecular 328 refractivity (31), and energy (15). Lastly, PC 4 considers the dissimilarity in: i) component of the dipole along the x axis (11), 329 ii) total dipole (14) and iii) minimum of the MEP surface (20). Regarding the samples belonging to the CC class, these 330 compounds were characterized by partner molecules with a similar behavior in terms of molecular complexity (21) and 331 octanol/water partition coefficient (10).



332

Figure 2. Loading plots related to the PCA decomposition. PC 1 vs. PC 2 (top-left), magnification of the PC 1 vs. PC 2 (top right), PC 2 vs. PC 3 (bottom-left), PC 3 vs. PC 4 (bottom-right). The fraction of variance explained by a given component is
 reported as a percentage value in parenthesis on the corresponding axis.

337 *4.1.2. Supervised pattern recognition*

The relationship between the class membership and the variables was exploited by means of PLS–DA. Four-Six LVs were retained according to maximum NERACC_% in cross validation. The PLS–DA model captured the 6372% and 6469% of the variance of the X_{cal} and Y_{cal} , respectively. The values of the predicted response \hat{y} in cross validation related to CC samples are plotted in Figure 2<u>S1</u>.



343 344 **Fie**

Figure 2: ŷ in cross validation for the CC samples. The dashed horizontal line shows the hard classification threshold of 0.50,
 halfway between the codified 0 (BM) and 1 (CC). Samples are marked according to their class (empty diamonds: cocrystals;
 black circles: binary mixtures).

A summary of the performance of the obtained model is reported in the confusion matrix (Table $\frac{12}{2}$), whereas a graphical representation of the estimated and predicted values \hat{y} for the CC class is reported in Figure S2. As reported in Table $\frac{12}{2}$, all the BM samples belonging to the test set were correctly classified except for one 5 samples, whereas only 5-3 out of 25-30 CC samples were wrongly assigned to the BM class, obtaining a NERACC_% of 8385%. Similarly, a high NERACC_% of 92% was obtained when the samples belonging to the calibration set were predicted by the model. The achieved results are extremely satisfactory, allowing for the *a priori* selection of the partner molecules required for the synthesis of novel cocrystals.

Table 12. Confusion matrix of the calibration and the test sets for the PLS–DA model. The second-last line shows the SEN_%
 for the modelled classes and last line shows the NERACC_% for the calibration and the test sets.

	Calibrat	ion set	Test set	
	Predicted as		Predicted	d as
	CC	BM	CC	BM
True CC	71<u>66</u>	5	20 27	<u>53</u>
True BM	7 <u>5</u>	<u>6351</u>	<u> 45</u>	9 19
SEN%	<u>93%</u>	<u>91%</u>	<u>90%</u>	<u>79%</u>

NERACC_% 92% 8385%

358

The distribution of the samples in the reduced space of the LVs can be observed by inspecting the score plot (Figure 3), whereas information regarding suspicious and/or influential samples can be retrieved by the squared residuals Q vs. Hotelling's T^2 and the residuals vs. leverage plots (Figure 4).

362



Figure 3. Distribution of the samples in the score space (LV 1 *vs.* LV 2). The fraction of variance explained by a given
 component is reported as a percentage value in parenthesis on the corresponding axis. Samples are marked according to their
 class (empty-magenta_diamonds: cocrystals; black-cyan_circles: binary mixtures).



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Figure 4. Reduced squared residuals Q vs. reduced Hotelling's T^2 plot, the dashed horizontal and vertical lines show the amplitude of the 95% confidence interval for both parameters (top). Residuals of CC samples vs. leverage plot, the dashed vertical line shows the leverage limit (bottom). Samples are marked according to their class (empty-magenta_diamonds: cocrystals; black-cyan circles: binary mixtures).

The maximum separation in the score space was provided by the first two LVs, with the CC samples <u>well groupedlocated</u> mainly at negative scores both on LV 1 and LV 2, especially in the III quadrant. By contrast, BM samples were more scattered and localized mostly on positive scores on LV 1.

A peculiar behavior was observed for samples (40) and (47): sample 40 showed Hotelling's T^2 -value outside the 95% confidence interval, whereas sample 47 showed both high Hotelling's T^2 and high squared residuals Q together with high leverage in the Y_{cal} -space. These CC samples were obtained by pairing fatty acids (lauric acid, 40, and palmitic acid, 47) with low molecular weight coformers, i.e., pyrazine and nicotinamide, respectively. The variables responsible for this behavior were related to the discrepancy in molecular dimensions between the two partner molecules.

- Another BM sample, i.e., the limonene/ascorbic acid (108), was characterized by high leverage and high Hotelling's *T*² value, due to different behavior in terms of hydrogen bond propensity and octanol/water partition coefficient. A similar behavior was observed also by two additional non influential BM samples based on ascorbic acid paired with cinnamaldehyde (89) and menthone (121).
- Finally, also the tartaric acid/pyrazine (63) and the adipic acid/hexamethyleneamine (12) CC samples showed Hotelling's T^2
- 388 values outside the 95% confidence interval and the latter resulted in having the highest residuals among CC samples.
- 389 Nevertheless, all these anomalous samples were correctly assigned in cross validation, and the variables with high contribution 390 (not shown) on their Hotelling's T^2 and high squared residuals O were characterized by low PLS weights.
- 391 On the other hand, one not anomalous CC sample and two not anomalous BM samples, namely carveol/isonicotinamide (20), 392 eugenol/pyrazine (103), geraniol/menthol (106), respectively, were characterized by very high residuals in absolute value, 393 thus being misclassified. This behavior can be ascribed to the fact that their features were inversely related to their respective 394 class. Also, the BM sample urea/hexamethylenetetramine (132) was misclassified: its features did not align with those of the 395 other samples; in fact, this sample hold high *Q* squared residuals. A peculiar behavior was observed for samples 27 and 29. B96 Sample 27 showed squared residual Q outside the 95% confidence interval, whereas sample 29 showed both high Hotelling's 397 T^2 and high squared residuals Q together with high leverage in the Y_{cal} space. These CC samples were obtained by pairing 398 fatty acids (lauric acid, 27, and palmitic acid, 29) with a low-molecular weight coformer, i.e., nicotinamide. The variables 399 responsible for this behavior can be related to the discrepancy in molecular dimensions between the two partner molecules.
- 400 The adipic acid/hexamethylenetetramine (11) CC sample was characterized by high leverage and high Hotelling's T^2 value,
- 401 <u>due to different behavior in terms of ruotable bonds and number of rings present in the structure. A similar behaviour, in terms</u>
- 402 of difference in the number of rings, was observed also by two additional non-influential BM samples based on
- 403 <u>hexamethylenetetramine paired with limonene (96) and menthone (103). In addition, these samples held high squared</u>
- 404 residuals *Q* and, therefore, were characterized by features that did not align with the ones of the other samples.
- Finally, the ferulic acid/pyrazine (22) CC sample showed Hotelling's T^2 values outside the 95% confidence interval due to the different behaviour of the partner molecules in terms of molecular weight, connectivity, surface area, and electronic properties, i.e., isotropic and anisotropic polarizability and number of valence electrons.
- 408 <u>Suspicious samples 22 (ferulic acid/pyrazine, CC) and 103 (menthone/hexamethylenetetramine, BM) appeared also to have</u>
- 409 high residuals in absolute value and were wrongly assigned to their class in cross validation. Along with them, also three not
- 410 anomalous CC samples and one not anomalous BM samples, namely cinnamaldehyde/4-hydroxybenzoic acid (17),

411 carvacrol/nicotinamide (60), thymol/tetramethylpyrazine (68) and eugenol/pyrazine (92), respectively, were misclassified. In 412 fact, their features were inversely related to their respective class. Nevertheless, the exclusion of the samples discussed above 413 from the calibration set would not have produced any difference in terms of rotation of the LV space due to their low leverage. 414 The correlation between class membership, coded in the Y_{cal} , and the predictors contained in the X_{cal} space can be observed 415 in the PLS weights plot (Figure S3). The variables involved in the discrimination are those whose weights follow the 416 discriminant direction: BM samples reach positive values of LV 1 and LV 2 in the Y_{est} loading space (not shown), whereas it 417 is the opposite for CC samples. Therefore, it can be inferred that significant differences in descriptors related to polarizability, 418 exposed surface, and volume, such as atom and bond count (2, 3), molecular volume (9), octanol-water partition coefficient (10), heteroatom count (27), topological polar surface area (22) are likely to prevent the formation of a cocrystal. 419 420 Information regarding the contribution of each predictor involved in the discrimination of the modelled classes can be inferred

421 by inspecting the pseudo-regression coefficients and the VIP score plots, reported in Figure 5. The latter parameter denotes 422 the relative importance of each predictor of the X_{cal} space in the PLS–DA model in explaining the class membership encoded in the Y_{cal} and may guide variable selection. Generally, a variable can be considered important with a VIP score > 1; by 423 424 contrast, a VIP score significantly lower than 1 indicates that a given variable is a good candidate for exclusion. According to 425 the negative sign of the pseudo-regression coefficients related to CC class, Therefore, it can be stated inferred that significant 426 differences in descriptors related to polarizability and exposed surface, such as atom and bond count (2, 3), octanol-water 427 partition coefficient (10), topological polar surface area (22) and heteroatom count (27) are likely to prevent the formation of 428 a cocrystal. It should be noted that this consideration agrees with what has emerged earlier from unsupervised modelling, and 429 it is largely in agreement with widely applied rules of thumb in crystal engineering. 430



431

432 Figure 5. Pseudo-regression coefficients plot for the CC class (top). VIP scores related to each variable included in the X_{ral} 433 space for the CC class, the significance threshold of 1 is depicted as a dashed horizontal line (bottom). Both plots report the 434 variable identification number on the *x* axis.

The variables that contributed most to the PLS weights were characterized by a VIP score > 1. AFinally, a reduced PLS–DA model based only on the important variables in agreement with the VIP scores was computed because of its ease of interpretation. A 4 LVs model was calculated according to the maximum NERACC_% in cross validation obtaining a classification performance very similar to that achieved by including all descriptors. The weights plot of the reduced model is provided in Figure 6, in which the weights of the variables involved in the discrimination are located on the positive and negative sides of the first LV.



Figure 6. PLS weights of the variables for LV 1 vs. LV 2 of the reduced model.

Despite the interpretation of the correlation pattern among the variables being not so straightforward, some main considerations can be drawn. As a general comment, it can be stated that a good balance in the hydrogen bond propensity between the two partner molecules has to be achieved, being the difference in hydrogen bond acceptors (7) and number of heteroatoms (27) on positive PLS weights on LV 1. In addition, the coformers should have a similar behavior in terms of polarity, being the dissimilarities in polar surface area (22) and octanol/water partition coefficient (10) on positive weights on LV 1. It should be noted that this consideration agrees with what has emerged earlier from unsupervised modelling, and it is largely in agreement with widely applied rules of thumb in crystal engineering.

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454 4.2. Prediction of unknown samples Benchmarking on an external validation set

455 In order to evaluate the predictive capability of the model, an external set of N = 58 binary combinations of partner molecules 456 was used. An overview of the involved samples is reported in Table S4 along with their estimated \hat{y} values.

Theis external validation set consists of 27 pairs classified experimentally as BMs and 31 pairs classified as CCs, i.e., 11 with our mechanochemistry/PXRD protocol and 20 retrieved from the CSD. The latter ones are reported in Section 3 of the

- 459 Supplementary Information together with their CSD refcode and the reference to the original publications.
- 460 <u>A graphical representation of the predicted values \hat{y} for the CC class is reported in Figure 7, as well as the squared residuals</u>
- 461 Q vs. Hotelling's T^2 plot. Although there were some samples that did not conform to the model space, not all of them have
- 462 <u>been systematically misclassified. The confusion matrix is reported in Table 3 to summarize the results.</u> In total, about 6274%
- 463 of the predictions were in agreement with the experimental results.

465 Table 3. Confusion matrix of the external validation set for the PLS–DA model. The second-last line shows the SEN_% for the

466 <u>modelled classes and last line shows the ACC_%.</u>

	External validation set		
	Predicted as		
	<u>CC</u>	BM	
Experimental CC	<u>29</u>	2	
Experimental BM	<u>13</u>	<u>14</u>	
<u>SEN</u> %	<u>94%</u>	<u>52%</u>	
ACC _%	<u>74%</u>		

467

468 Specifically, 26-29 CC cases out of 31 were correctly classified. On the other hand, there were 17-13 false positive results and 469 10-14 cases in which the pairs were correctly identified as BMs. The model appears therefore to be quite conservative in 470 discarding the possibility of cocrystallization; hence, fewer potential new materials could be overlooked. Similar behavior 471 was observed also in the test set, meaning that the missed discovery rate does not get worse when working on completely external data, thus demonstrating the stability of the model. In addition, the fraction of misclassified CC samples (see Table 472 473 2 and Table 3) does not change significantly in the external validation set with respect to the test set. 474 As shown in Figure 7, Table 3 and Table S4, only 2 pairs of partner molecules predicted as BMs actually formed CCs, thus 475 producing 2 false negative results. It can be stated that these Moreover, all the 5 false negative results were borderline cases 476 when the estimated error [63] on the prediction was taken into consideration (data not shown): in fact their $\hat{\mathbf{p}}$ value was close 477 to the classification threshold of 0.53. The same is true for $\frac{8}{6}$ of the false positive cases, while in the remaining $\frac{9}{7}$ cases the 478 model confidently classified the pairs as CCs, in contrast to our experimental results. These findings could be ascribed to the 479 use of different preparation methods other than mechanochemical grinding. This hypothesis is somehow supported by the fact 480 that 19 out of the 20 CCs retrieved from the CSD, then prepared with a variety of methods, were correctly identified.



Figure 7. \hat{y} (CC samples) for the external validation set, the dashed horizontal line shows the hard classification threshold of 0.53 (left). Reduced squared residuals *Q vs.* reduced Hotelling's T^2 plot, the dashed horizontal and vertical lines show the amplitude of the 95% confidence interval for both parameters (right). Samples are marked according to their class (magenta diamonds: cocrystals; cyan circles: binary mixtures).

482

488 **5.** Conclusion

This study highlighted the ability of a simple PLS–DA model to predict cocrystal formation without any *a priori* knowledge of the specific role of the involved partner molecules. Information deriving from both successful and unsuccessful cocrystallization experiments was used. The major advantage of the proposed methodology relies on the reduction of the experimental effort required for both the synthesis and characterization of new crystalline structures.

- The model allows us to predict cocrystallization propensity with a $\frac{6274}{5}\%$ of the predictions in agreement with the experimental results. Considering that the model was obtained on a training set spanning different molecular characteristics,
- it can be stated that it is suitable for a fairly general applicability.
- Indeed, once in possess of the set of chemical descriptors for the molecules of interest, it is sufficient to calculate the absolute value of their difference and perform a linear combination using the pseudo-regression coefficients to obtain a prediction on cocrystal formation. The precalculated values for the set of descriptors comprising 2193 GRAS molecules are available in the Supplementary Material. By applying the proposed methodology, seven ten new cocrystals were discovered and <u>an</u> additional four compounds waswere obtained by chance.

Another figure of merit of the proposed approach is the possibility of understanding through the inspection of PLS weights how the degree of similarity in terms of molecular features of the two partner molecules is correlated with the possibility of obtaining a cocrystal.

On a closing note, we would like to strongly encourage scientists to report failed attempts at cocrystallization along with the technique used, as access to this information could play a pivotal role in refining predictive models, making them less sensitive to selective reporting bias.

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515 Authorship Contribution Statement

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- 517 (equal), Writing Original Draft Preparation (lead), Writing Review & Editing (equal).
- 518 Fabio Montisci: Investigation (equal), Data Curation (lead), Validation (equal), Visualization (equal), Writing Original
- 519 Draft Preparation (supporting), Writing Review & Editing (equal).
- 520 Federica Bianchi: Conceptualization (equal), Methodology (lead), Resources (equal), Supervision (equal), Writing Review
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- 524 **Claudia Carraro:** Investigation (equal), Data Curation (equal).
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- 533 Review & Editing (equal).
- 534
- 535

536 **Declaration of Competing Interest**

- 537 The authors declare that they have no known competing financial interests or personal relationships that could have appeared
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Highlights

A QSPR model for the discovery of cocrystals made by EOs and other GRAS molecules Training set based on failed and successful cocrystallization experiments Correct classification rate of 85% on the test set Broad applicability and reduced experimental effort

Declaration of interests

⊠The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

□The authors declare the following financial interests/personal relationships which may be considered as potential competing interests:

Author Statement

Fabio Fornari: Conceptualization (equal), Formal Analysis (lead), Methodology (equal), Validation (equal), Visualization (equal), Writing – Original Draft Preparation (lead), Writing – Review & Editing (equal).

Fabio Montisci: Investigation (equal), Data Curation (lead), Validation (equal), Visualization (equal), Writing – Original Draft Preparation (supporting), Writing – Review & Editing (equal).

Federica Bianchi: Conceptualization (equal), Methodology (lead), Resources (equal), Supervision (equal), Writing – Review & Editing (equal).

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Alessia Bacchi: Conceptualization (lead), Funding Acquisition (equal), Resources (equal), Supervision (equal), Writing – Review & Editing (equal).

Supplementary Material

Click here to access/download Supplementary Material Supplementary Material.docx Supplementary Material

Click here to access/download Supplementary Material GRAS_molecules_dataset.xlsx Supplementary Material

Click here to access/download Supplementary Material PLSDAmodel.m

1	Chemometric-Assisted Cocrystallization: Supervised Pattern Recognition for Predicting the
2	Formation of New Functional Cocrystals
3	
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27 Abstract

Owing to the antimicrobial and insecticide properties, the use of natural compounds like essential oils and their active components has proven to be an effective alternative to synthetic chemicals in different fields ranging from drug delivery to agriculture and from nutrition to food preservation. Their limited application due to the high volatility and scarce water solubility can be expanded by using crystal engineering approaches to tune some properties of the active molecule by combining it with a suitable partner molecule (coformer). However, the selection of coformers and the experimental effort required for discovering cocrystals are the bottleneck of cocrystal engineering. This study explores the use of chemometrics to aid the discovery of cocrystals of active ingredients suitable for various applications. Partial Least Squares-Discriminant Analysis is used to discern cocrystals from binary mixtures based on the molecular features of the coformers. For the first time, by including failed cocrystallization data and considering a variety of chemically diverse compounds, the proposed method resulted in a successful prediction rate of 85% for the test set in the model validation phase and of 74% for the external test set.

50 Keywords

51 cocrystal, crystal engineering, chemoinformatics, chemometrics, partial least square discriminant analysis, Quantitative

52 Structure–Property Relationship

53	List of abbreviations (sorted alphabetically)
54	ACC _% ; Accuracy
55	BM: Binary Mixture
56	CC: Cocrystal
57	CCDC: Cambridge Crystallographic Data Centre
58	CSD: Cambridge Structural Database
59	EO: Essential Oil
60	FDA: Food and Drug Administration
61	GRAS: Generally Recognized As Safe
62	LV: Latent Variable
63	MEP: Molecular Electrostatic Potential
64	PC: Principal Component
65	PCA: Principal Component Analysis
66	PLS-DA: Partial Least Squares-Discriminant Analysis
67	PXRD: Powder X-Ray Diffraction
68	QSPR: Quantitative Structure–Property Relationship
69	SEN _% : Sensitivity
70	VIP: Variable Importance in Projection.
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81 **1. Introduction**

In the last few decades, the use of agrochemicals and food preservatives has grown exponentially as a direct consequence of the rapid increase of the world population [1,2]. Owing to their potential adverse effects on both human health and environment [2–4], alternative strategies based on the use of more sustainable chemicals have been proposed to support the food system. Being able to exert antimicrobial, insecticidal, and antioxidant properties [5,6], essential oils (EOs) and their active components have been used as green substitutes of synthetic chemicals to extend the shelf-life of foodstuff and in pests control [7–9]. These compounds are Generally Recognized As Safe (GRAS) by the Food and Drug Administration (FDA) [10], however, despite their appealing properties, their use is limited by their high volatility and poor stability [7,9,11,12].

In fact, physicochemical properties of materials play a key role in determining whether a chemical is suitable for a specific purpose, thus strongly affecting its field of application. Scientists have always desired to obtain materials with target properties, and crystal engineering is one of the most interesting approaches to synthesize a great variety of crystalline materials for applications in various fields, ranging from pharmaceuticals to agrochemicals, and from nutraceuticals to cosmetics [13–16]. The basic idea of crystal engineering is related to the possibility of controlling the crystal structure of molecules and, therefore, the properties of the resulting solids. Polymorphism, vitrification and cocrystallization are some of the available strategies to modify the intrinsic properties of molecules without the need of synthetic modifications [17–21].

96 Cocrystals are multicomponent crystalline solid materials in which the constituents (i.e., coformers) are bound in a well-97 defined stoichiometric ratio [22,23] *via* non-covalent interactions (e.g., hydrogen bonds, halogen bonds, π - π stacking) within 98 the same crystal structure. Cocrystallization allows for the combination of the desired molecule of interest with properly 99 selected partner molecules, paving the way to an array of potential materials with enhanced properties [19,24,25]. Within this 100 frame of reference, cocrystals based on the active components of EOs have been proposed as active ingredients for food 101 packaging, agrochemical and pharmaceutical applications [7,19,26,27].

Despite the great advantages offered by cocrystallization, the proper degree of complementarity between the two partner molecules required to obtain crystalline materials with the desired properties is not easy to assess [28–31]. In this context, the selection of coformers and the great effort required for both the systematic experimental screening and careful characterization of the products derived from the combination of all the possible coformer pairs represent the major bottleneck of cocrystal engineering. Computational techniques represent a powerful tool to reduce the experimental effort required for the discovery of new cocrystals, enabling to evaluate beforehand whether a cocrystal can be obtained starting from pre-selected coformers. These *in silico* strategies can be based on the calculation of a variety of parameters useful to predict the formation of a 109 cocrystal, such as lattice energy [32], solubility [33], hydrogen bond propensity along with the quantitation of molecular
110 interaction energy [29,30], and molecular complementarity [34].

Despite the massive efforts spent to develop a method to predict cocrystal formation, at present none of the proposed strategieshas proven to be both totally reliable and easy to apply.

Chemometrics could play a pivotal role in cocrystal discovery: up to now only few Machine Learning methods have been 113 114 proposed in predicting cocrystal formation, enabling the screening of new cocrystals once a supervised model is properly 115 trained and validated. In the study proposed by Devogelaer et al., information of successful cocrystallization experiments was 116 directly taken from the Cambridge Structural Database (CSD) [35] and Artificial Neural Networks (ANN) were used to predict 117 the formation of new cocrystals [36]. Similarly, Wang et al. relied on a consensus method based on multiple Random Forest 118 algorithms, in which the successful cocrystallization dataset was integrated with randomly generated failed cocrystallization 119 data [37]. These approaches are reported in the literature as network-based methods. Additional studies were based on the use 120 of successful and unsuccessful cocrystallization datasets obtained from experimentation, literature, and/or the CSD for 121 screening specific classes of coformers. Within this framework, Przybyłek et al. used Multivariate Adaptive Regression 122 Splines to predict the formation of dicarboxylic and phenolic acid-based cocrystals [38,39], whereas Wicker et al. focused on 123 variously substituted benzoic acids and benzamides using a Support Vector Machine algorithm [24]. Vriza et al. used an 124 ensemble one-class classification method to aid the discovery of $\pi - \pi$ cocrystals, thus giving a great contribution in enriching 125 one of the most under-represented classes of cocrystals in the CSD [40]. Most recently, Mswahili et al. developed a cocrystal 126 screening method based on ANN by using both successful and unsuccessful experimental cocrystallization data retrieved from 127 the literature and a plethora of molecular descriptors calculated using Mordred [41,42].

In the frame of a research activity dealing with the synthesis of new functional cocrystals based on the active constituents of EOs and other GRAS molecules to broaden their applicability in the industrial field [7,19], we propose a chemometric approach to aid the discovery of new cocrystalline materials.

For the first time, a training set based on the results of failed (binary mixtures, BM) and successful (cocrystal, CC) cocrystallization experiments was used for the computation of a Quantitative Structure–Property Relationship-like (QSPR) model based on Partial Least Squares–Discriminant Analysis (PLS–DA), after preliminary exploratory analysis by Principal Component Analysis (PCA). The PLS-DA approach, with respect to network-based methods offers the advantages of having only one parameter to optimize, i.e., the model dimensionality, and the direct interpretation of the importance of descriptors in classification, while highlighting their interplay (by inspection of weights and loadings plots).

- 137 The effectiveness of the study relies on the use of compounds belonging to different chemical classes and a reduced number
- 138 of 1D, 2D, and 3D molecular descriptors of various nature (e.g., constitutional, geometric, physical, topological, and surface
- 139 area-based descriptors) [24,38,39,43], enabling the high-throughput screening of novel cocrystalline materials and offering
- 140 maximum flexibility and effectiveness at a minimum computational and experimental cost.
- 141

142 2. Experimental Procedures

143 2.1. Mechanochemical protocol and class assignation

All the molecules in the dataset were chosen among the list of GRAS molecules drawn up by the FDA [10]. Selected pair of molecules among the chosen ones were assigned either to the CC class or to the BM class. Pairs of molecules in the dataset for which a cocrystal structure was already described in literature were individuated in the Cambridge Structural Database (CSD) [35] with the Cambridge Crystallographic Data Centre (CCDC) software ConQuest [44] and visualized with Mercury [45]. They are reported in Section 3 of the Supplementary Material, together with their unique CSD refcode and the reference to the original publications.

150 Cocrystallization for all the pairs with no known structure in literature was instead attempted with the following mechanochemical protocol. All the reagents employed were commercially available and used as such in all the experiments. 151 152 Equimolar amounts of each reagent were directly mixed in an agate mortar and subjected to manual grinding for 10-15 153 minutes, without using any solvent. The resulting powder samples were collected in closed vials. Assignation to CC or BM 154 classes was performed by comparing the Powder X-ray Diffraction (PXRD) pattern of the ground sample with those of the 155 pure reagents. Possible occurrence of polymorphic transitions for the reagents was excluded by comparing the experimental 156 PXRD data after milling with the calculated pattern of all the known crystalline forms of the reagents. The occurrence of new peaks, unexplained by the presence of unreacted reagents, was taken as indication that cocrystallization had occurred and the 157 158 sample was assigned to the CC class. In case no additional peaks appeared in the PXRD pattern the sample was instead 159 classified as a BM.

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162 2.2. Powder X-ray diffraction

Typically, PXRD data were collected on a Rigaku Smartlab XE diffractometer in θ - θ Bragg-Brentano geometry with Cu Kα radiation. The samples were placed on glass supports and exposed to radiation $(1.5^\circ \le 2\theta \le 50^\circ)$ at a scan rate of 10°/min. The diffracted beam was collected on a 2D Hypix 3000 solid state detector. 5° radiant soller were used as a compromise for high flux and moderate peak asymmetry at low angles. Beam stopper and anti-scatterer air component were used to mitigate the profile at low angle. In some rare cases, the data were collected on a Thermo Fisher Scientific ARL X'TRA diffractometer in θ - θ Bragg-Brentano geometry with Cu Kα radiation (3° $\le 2\theta \le 30^\circ$ at a scan rate of 5°/min, or 3° $\le 2\theta \le 40^\circ$ at a scan rate of 0.3°/min).

170

171 **3.** Computational Methods

172 3.1. Molecular descriptors calculation

For each molecule 31 molecular descriptors were calculated (Table S1). A theoretical background for the less knowndescriptors is given in Section 4 of the Supplementary Material.

175 The molecular weight, the number of atoms, the number of bonds, the number of hydrogen bond donor sites and the number 176 of hydrogen bond acceptor sites were calculated with FLAP software (Fingerprint for Ligand and Protein) [46] at pH 7.0, 177 using the 3D structures of all molecules in SDF format as input (downloaded from the PubChem database). The number of 178 rotatable bonds, the number of rings, the hydrophobicity (accounted as the number of hydrophobic centers), the logP 179 (logarithm of octanol/water partition coefficient), the molecular volume, the total molecular dipole moment (based on point 180 charge distribution in the molecule), and its components along the axes (using the principal axes of the molecular graph) were 181 then calculated for the same structures using Sybyl 8.1 [47] (www.tripos.com) and taking in consideration the protonation 182 state of molecules. The same software was also used to estimate the strain energy of the molecule without performing any geometry optimization. This energy term relies on an electrostatic calculation from atomic charges using the internal Tripos 183 184 force field [48]. For the estimation of molecular volume and dipole moment, a specific SPL script was employed. The 185 calculated volume is enclosed in a water-accessible surface computed at a repulsive interaction energy of 0.20 kcal/mol with 186 a water probe. A custom Python script was used to automatically calculate the Solvent Accessible Surface Area (SASA) in 187 PyMol 2.0 [49], with the dot density parameter set to 4. The number of heteroatoms, the number of valence electrons, and the indexes ${}^{0}\chi$, ${}^{0}\chi^{\nu}$, ${}^{0}\chi^{\nu}$, ${}^{\alpha}\mu_{K}$, ${}^{1}\kappa_{\alpha}$, LabuteASA, SMR_VSA, PEOE_VSA, and TPSA were obtained running a Python 3.7 code with 188 189 the open-source cheminformatics toolkit RDKit Q4 2013 [50]. The average isotropic polarizability α_{iso} , the polarizability anisotropy $\Delta \alpha$, and the Molecular Electrostatic Potential (MEP) were calculated with Gaussian 16 [51] following the *in vacuo*

191 Density-Functional Theory optimization of all the molecules, employing the hybrid functional B3LYP and the People double-

192 z basis set 6-31+g(d,p).

193 Postprocessing of the MEP to extract critical points at a given electron density isosurface was performed with a custom Python 194 3.6.1 script on a three-dimensional map (cube format) with a sampling density of 6 points/Bohr along the three directions. 195 The MEP was analyzed at an electron density isosurface of 0.002 a.u. with a tolerance of 0.001 a.u., meaning that only MEP 196 values corresponding to regions of space with electron density in the 0.001–0.003 a.u. range were considered. A first set of 197 critical points was identified comparing MEP values of each cube point with those of its 6 nearest neighbors. A point was considered a local minimum if the number of nearest neighbors with higher MEP was greater or equal to a given integer (4). 198 199 Likewise, a point was considered a local maximum if the number of nearest neighbors with lower MEP was greater or equal 200 to the same integer. This first step yielded a large number of candidate critical points encompassing a wide range of MEP 201 values. Since our focus was on identifying the regions of the molecules likely to be involved in strong hydrogen bonds within 202 the cocrystal, in a second step this first set of points was filtered based on the MEP values of the global minimum and 203 maximum. This was done as follows: for each local minimum (maximum), the ratio between its MEP value and that of the 204 global minimum (maximum) was computed, and the point was kept only if the ratio exceeded a given threshold (0.1). In this 205 way, only points corresponding to shallow critical points were discarded. This step allowed to identify the MEP isosurface 206 regions corresponding to hydrogen bond donors and acceptors. However, due to the rugged character of the MEP map. 207 multiple critical points of the same type could appear in close proximity. To univocally map a given region of the isosurface 208 to a MEP value, critical points close to each other (below a distance threshold of 1.0 Bohr) were merged iteratively, keeping 209 only the lower MEP point for minima and higher MEP point for maxima. The algorithm then provided the final set of MEP 210 critical points at the given electron density isosurface.

211

212 3.2. Data analysis

The entire data analysis was carried out in MATLAB R2019a environment (Mathworks, Natick, Massachusetts, USA) with
the aid of the PLS_Toolbox 8.7.1 (Eigenvector Research Inc., Washington, USA) chemometric package.

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- 216

217 *3.2.1. Data preprocessing*

Each sample was described by m = 31 variables (Table S1): the absolute value of the difference between the molecular descriptors of the two partner molecules was calculated, thus obtaining the predictor matrix X (181 × 31). The class membership was binary encoded (1: belonging to the class; 0: otherwise) in a dummy matrix Y (181 × 2) with each column representing one of the two modelled classes. The dataset was split in two subsets by using the Duplex algorithm [52]: 70% of the data were used as calibration set, X_{cal} (127 × 31) and Y_{cal} (127 × 2), whereas the remaining 30% were used as test set,

- 223 X_{test} (54 × 31) and Y_{test} (54 × 2). The calibration set and the test set are reported in Table S2 and Table S3, respectively.

Before carrying out both exploratory multivariate data analysis and the computation of the supervised model, the calibration matrix X_{cal} was preprocessed column-wise by performing mean centering and scaling to unit variance. Mean centering was applied on the response matrix Y_{cal} to ensure the stability of the model.

227

228 3.2.2. Exploratory multivariate data analysis

PCA [53–55] was carried out preliminarily on the calibration set X_{cal} to assess the distribution of the samples and to check for potential data structures. Reduction of data dimensionality is carried out through the linear combination of the original variables in a set of orthogonal ones, i.e., Principal Components (PCs), which identify the direction of maximum variance. This is summarized in the decomposition equation:

 $X_{\rm cal} = TP^{\rm T} + E$

where T and P represent, respectively, the coordinates of the samples projected in the reduced space, i.e., the scores, and the weights each original variable has on a given PC, i.e., the loadings. The deviations from the model are accounted in the error matrix E.

237

238 *3.2.3. Supervised pattern recognition*

PLS-DA [56,57] was used to discriminate pairs of partner molecules whose combination forms CCs from the ones giving BMs. PLS-DA is based on PLS regression [58]. Briefly, this supervised technique decomposes the predictor matrix X_{cal} and the dependent variables matrix Y_{cal} in a PCA-like way and imposes inner linear relationships between the X_{cal} and the Y_{cal} scores as follows:

where T and U, are the X_{cal} and the Y_{cal} scores, respectively. This is accomplished by rotating the Latent Variable (LV) space of X_{cal} through a weight matrix W in a way that maximizes the covariance between T and U. The PLS regression model is summarized as:

247
$$Y_{cal} = X_{cal}B + E$$

248 where *E* is the error matrix and *B* is the pseudo-regression coefficient matrix expressed according to the following equation:

$$B = W(P^T W)^{-1} \operatorname{diag}(b) Q$$

250 where P and Q are the X_{cal} and the Y_{cal} loadings, respectively.

In this case, the dependent variables in the Y_{cal} matrix are defined as dummy variables, one for each modelled class, taking values of 1 if the sample belongs to the class and 0 otherwise. Current implementation of PLS–DA may differ on the basis of how the classification rule is defined. In this work, a pure discriminant rule (samples are assigned univocally to only one category) was applied, and thus a sample is assigned to the class for which the predicted response \hat{y} is the highest (i.e., \hat{Y} values are continuous and not dummy as they were codified).

The proper number of LVs was chosen according to the maximum accuracy (ACC_%; i.e., the percentage of samples correctly assigned to the respective class) in leave-more-out cross validation, adopting a Venetian blinds cancellation scheme with 10 splits (blind thickness: 1). This operation was carried out by running a custom MATLAB routine. The performance of the classification model was evaluated both on the calibration and the test sets in terms of ACC_% as well as showing the confusion matrix. In addition, the sensitivity (SEN_%, i.e., the percentage of samples within a class that were correctly assigned to their class) was calculated for both classes.

The importance of each predictor was estimated in terms of Variable Importance in Projection (VIP score) [56]. The VIP score of the jth variable in the *X* space is defined as the component-wise sum of its PLS weight w_j on the fth component multiplied by the fraction of variance of the *Y* explained by that component, according to the following equation:

265
$$VIP_j^2 = \frac{1}{SSY_{\text{tot}}F} \sum_{f=1}^F w_{jf}^2 SSY_f J$$

266 where *J* is the number of variables in the *X* space and *F* is the number of LVs that were retained. Since:

$$\sum_{j=1}^{J} VIP_j^2 = J$$

the proposed threshold for determining whether a variable could be considered important is set to 1..

- 269 Finally, the predictive capability of the model was evaluated on an external set of N = 58 binary combinations of partner
- 270 molecules. An overview of the involved samples is reported in Table S4 along with their estimated \hat{y} values.
- 271 The number of entries in the BM and CC classes for training, test, and external validation sets is reported in Table 1.
- 272
- 273 Table 1. Number of CC and BM samples in the calibration, test, and external validation set. The last column reports the total

274 number *N* of samples *per* set.

	CC	BM	N
Calibration set	71	56	127
Test set	30	24	54
External validation set	31	27	58

276 **4. Results and Discussion**

The cocrystallization experiments were carried out mechanochemically by manual neat grinding of the two substances. This method was selected among many possible others due to its simplicity and promptness, allowing us to screen several molecular pairs in a standardized way [21,59]. The classification as BM or CC was assessed by PXRD patterns (available in Section 3 of the Supplementary Material). Cocrystals already present in the CSD were also included in our dataset.

281

282 4.1. Data analysis

283 Data handling prior to analysis can affect the way the model is trained, thus having consequences on its interpretation.

284 Since samples are the result of the combination of two partner molecules, each one described by its own set of descriptors, 285 the concatenation strategy, i.e., listing the descriptors of the first coformer followed by those of the second coformer, was 286 discarded due to the lack of commutation between the two sets of descriptors. In fact, the order in which the molecular 287 descriptors are listed represents an *a priori* decision on which compound is acting as molecule of interest or partner molecule. 288 In our case this would be sub-optimal since many of the molecules in our dataset could assume both roles. Therefore, in order 289 to address the problem described above a commutative strategy capable of avoiding the production of indeterminate forms 290 should be chosen. Considering that in the dataset many constitutional molecular descriptors were characterized by a few non-291 zero values, the calculation of both products and ratios between the descriptors was discarded since additional zero values or 292 indeterminate forms could be generated. Also, the division is non-commutative.

In order to combine the two partner molecules without imposition on their role, the absolute value of the difference between the molecular descriptors of the partner molecules was calculated and used to describe each binary combination [38,39], giving maximum flexibility to the model. The information achieved is still relevant and easily interpretable since it is related to the dissimilarity between the descriptors. In fact, absolute value of the differences between descriptors for each case are the elements of the Manhattan distance [60], one of the possible indexes to account for the dissimilarity between cases in a multivariate way.

In the present study, the use of basic linear modelling methods was preferred with respect to non-linear modelling, such as
ANN [61,62], to keep the calculations as simple as possible, and to ensure a certain degree of interpretability of the results.
Furthermore, the number of samples is too limited to ensure proper tuning of the ANN hyperparameters.

302 For a fruitful discussion, samples and variables are reported in the text according to their identification number as follows: i)

- samples are written in plain text; ii) variables are underlined. The key is available in Tables S1–S3.
- 304

305 *4.1.1. Exploratory multivariate data analysis*

After data preprocessing, PCA was used in an exploratory way to assess the presence of potential data structures in thecalibration set.

Four PCs were retained explaining 67% of the variance. As shown in the score plots (Figure 1), a mild separation was present in the PC 2 *vs.* PC 3 score plot, with the groups separated by the bisecting line of the II and the IV quadrant. In addition, most of the CC samples occupied the III quadrant of the PC 2 *vs.* PC 3 score plot and were, in general, less scattered than BM samples.

312





318 As for the loading plots depicted in Figure 2, it can be observed that PC 1 explains the features related both to differences in 319 molecular dimensions (e.g., <u>1</u>, <u>2</u>, <u>3</u>, <u>9</u>, <u>16</u>, <u>29</u>) and connectivity (<u>23</u>, <u>24</u>, <u>25</u>, <u>26</u>). PC 2 considers the dissimilarities in the 320 electronic properties (e.g., <u>6</u>, <u>7</u>, <u>19</u>, <u>22</u>, <u>30</u>) of the two partner molecules as well as the difference in their number of 321 heteroatoms (27). PC 3 accounts for more specific features, such as differences in molecular complexity (21), molecular 322 refractivity (31), and energy (15). Lastly, PC 4 considers the dissimilarity in: i) component of the dipole along the x axis (11), 323 ii) total dipole ($\underline{14}$) and iii) minimum of the MEP surface ($\underline{20}$). Regarding the samples belonging to the CC class, these 324 compounds were characterized by partner molecules with a similar behavior in terms of molecular complexity (21) and 325 octanol/water partition coefficient (10).



- Figure 2. Loading plots related to the PCA decomposition. PC 1 *vs.* PC 2 (top-left), magnification of the PC 1 *vs.* PC 2 (topright), PC 2 *vs.* PC 3 (bottom-left), PC 3 *vs.* PC 4 (bottom-right). The fraction of variance explained by a given component is
 reported as a percentage value in parenthesis on the corresponding axis.
- 330
- 331 *4.1.2. Supervised pattern recognition*

The relationship between the class membership and the variables was exploited by means of PLS–DA. Six LVs were retained according to maximum ACC_% in cross validation. The PLS–DA model captured the 72% and 69% of the variance of the X_{cal} and Y_{cal} , respectively. The values of the predicted response \hat{y} in cross validation related to CC samples are plotted in Figure S1.

- A summary of the performance of the obtained model is reported in the confusion matrix (Table 2), whereas a graphical representation of the estimated and predicted values \hat{y} for the CC class is reported in Figure S2. As reported in Table 2, all the BM samples belonging to the test set were correctly classified except for 5 samples, whereas only 3 out of 30 CC samples were wrongly assigned to the BM class, obtaining a ACC_% of 85%. Similarly, a high ACC_% of 92% was obtained when the samples belonging to the calibration set were predicted by the model. The achieved results are extremely satisfactory, allowing for the *a priori* selection of the partner molecules required for the synthesis of novel cocrystals.
- 342

Table 2. Confusion matrix of the calibration and the test sets for the PLS–DA model. The second-last line shows the SEN[%]
for the modelled classes and last line shows the ACC[%] for the calibration and the test sets.

	Calibratio	n set	Test set	
	Predicted as		Predicted a	IS
	CC	BM	CC	BM
True CC	66	5	27	3
True BM	5	51	5	19
SEN%	93%	91%	90%	79%
ACC _%	92%		85%	

The distribution of the samples in the reduced space of the LVs can be observed by inspecting the score plot (Figure 3), whereas information regarding suspicious and/or influential samples can be retrieved by the squared residuals Q vs. Hotelling's T^2 and the residuals vs. leverage plots (Figure 4).

349



350

Figure 3. Distribution of the samples in the score space (LV 1 *vs.* LV 2). The fraction of variance explained by a given
component is reported as a percentage value in parenthesis on the corresponding axis. Samples are marked according to their
class (magenta diamonds: cocrystals; cyan circles: binary mixtures).





Figure 4. Reduced squared residuals Q vs. reduced Hotelling's T^2 plot, the dashed horizontal and vertical lines show the amplitude of the 95% confidence interval for both parameters (top). Residuals of CC samples vs. leverage plot, the dashed vertical line shows the leverage limit (bottom). Samples are marked according to their class (magenta diamonds: cocrystals; cyan circles: binary mixtures).



scores on LV 1. By contrast, BM samples were more scattered and localized mostly on positive scores on LV 1.

A peculiar behavior was observed for samples 27 and 29. Sample 27 showed squared residual Q outside the 95% confidence interval, whereas sample 29 showed both high Hotelling's T^2 and high squared residuals Q together with high leverage in the *Y*_{cal} space. These CC samples were obtained by pairing fatty acids (lauric acid, 27, and palmitic acid, 29) with a low-molecular

- 366 weight coformer, i.e., nicotinamide. The variables responsible for this behavior can be related to the discrepancy in molecular
- 367 dimensions between the two partner molecules.

The adipic acid/hexamethylenetetramine (11) CC sample was characterized by high leverage and high Hotelling's T^2 value, due to different behavior in terms of ruotable bonds and number of rings present in the structure. A similar behaviour, in terms of difference in the number of rings, was observed also by two additional non-influential BM samples based on hexamethylenetetramine paired with limonene (96) and menthone (103). In addition, these samples held high squared residuals Q and, therefore, were characterized by features that did not align with the ones of the other samples.

Finally, the ferulic acid/pyrazine (22) CC sample showed Hotelling's T^2 values outside the 95% confidence interval due to the different behaviour of the partner molecules in terms of molecular weight, connectivity, surface area, and electronic properties, i.e., isotropic and anisotropic polarizability and number of valence electrons.

Suspicious samples 22 (ferulic acid/pyrazine, CC) and 103 (menthone/hexamethylenetetramine, BM) appeared also to have 376 377 high residuals in absolute value and were wrongly assigned to their class in cross validation. Along with them, also three not 378 anomalous CC samples and one not anomalous BM samples, namely cinnamaldehyde/4-hydroxybenzoic acid (17), 379 carvacrol/nicotinamide (60), thymol/tetramethylpyrazine (68) and eugenol/pyrazine (92), respectively, were misclassified. In 380 fact, their features were inversely related to their respective class. Nevertheless, the exclusion of the samples discussed above 381 from the calibration set would not have produced any difference in terms of rotation of the LV space due to their low leverage. 382 The correlation between class membership, coded in the Y_{cal} , and the predictors contained in the X_{cal} space can be observed 383 in the PLS weights plot (Figure S3).

384 Information regarding the contribution of each predictor involved in the discrimination of the modelled classes can be inferred 385 by inspecting the pseudo-regression coefficients and the VIP score plots, reported in Figure 5. The latter parameter denotes 386 the relative importance of each predictor of the X_{cal} space in the PLS–DA model in explaining the class membership encoded 387 in the Y_{cal} and may guide variable selection. Generally, a variable can be considered important with a VIP score > 1; by 388 contrast, a VIP score significantly lower than 1 indicates that a given variable is a good candidate for exclusion. According to 389 the negative sign of the pseudo-regression coefficients related to CC class, it can be stated that significant differences in 390 descriptors related to polarizability and exposed surface, such as atom and bond count (2, 3), octanol-water partition 391 coefficient (10), topological polar surface area (22) and heteroatom count (27) are likely to prevent the formation of a 392 cocrystal. It should be noted that this consideration agrees with what has emerged earlier from unsupervised modelling, and 393 it is largely in agreement with widely applied rules of thumb in crystal engineering.



Figure 5. Pseudo-regression coefficients plot for the CC class (top). VIP scores related to each variable included in the X_{cal} space for the CC class, the significance threshold of 1 is depicted as a dashed horizontal line (bottom). Both plots report the variable identification number on the *x* axis.

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Finally, a reduced PLS–DA model based only on the important variables in agreement with the VIP scores was computed because of its ease of interpretation. A 4 LVs model was calculated according to the maximum ACC_% in cross validation obtaining a classification performance very similar to that achieved by including all descriptors. The weights plot of the reduced model is provided in Figure 6, in which the weights of the variables involved in the discrimination are located on the positive and negative sides of the first LV.



406 Figure 6. PLS weights of the variables for LV 1 vs. LV 2 of the reduced model.

405

408 4.2. Benchmarking on an external validation set

The external validation set consists of 27 pairs classified experimentally as BMs and 31 pairs classified as CCs, i.e., 11 with our mechanochemistry/PXRD protocol and 20 retrieved from the CSD. The latter ones are reported in Section 3 of the

411 Supplementary Information together with their CSD refcode and the reference to the original publications.

412 A graphical representation of the predicted values \hat{y} for the CC class is reported in Figure 7, as well as the squared residuals

413 Q vs. Hotelling's T^2 plot. Although there were some samples that did not conform to the model space, not all of them have

414 been systematically misclassified. The confusion matrix is reported in Table 3 to summarize the results. In total, about 74%

- 415 of the predictions were in agreement with the experimental results.
- 416

417 Table 3. Confusion matrix of the external validation set for the PLS–DA model. The second-last line shows the SEN_% for the

418 modelled classes and last line shows the ACC_%.

	External validation set Predicted as	
	CC	BM
Experimental CC	29	2
Experimental BM	13	14
SEN _%	94%	52%

ACC _%	74%	



420 Specifically, 29 CC cases out of 31 were correctly classified. On the other hand, there were 13 false positive results and 14 421 cases in which the pairs were correctly identified as BMs. The model appears therefore to be quite conservative in discarding 422 the possibility of cocrystallization; hence, fewer potential new materials could be overlooked. In addition, the fraction of 423 misclassified CC samples (see Table 2 and Table 3) does not change significantly in the external validation set with respect 424 to the test set.

As shown in Figure 7, Table 3 and Table S4, only 2 pairs of partner molecules predicted as BMs actually formed CCs, thus producing 2 false negative results. It can be stated that these results were borderline cases when the estimated error [63] on the prediction was taken into consideration (data not shown): in fact their \hat{y} value was close to the classification threshold of 0.53. The same is true for 6 of the false positive cases, while in the remaining 7 cases the model confidently classified the pairs as CCs, in contrast to our experimental results. These findings could be ascribed to the use of different preparation methods other than mechanochemical grinding. This hypothesis is somehow supported by the fact that 19 out of the 20 CCs retrieved from the CSD, then prepared with a variety of methods, were correctly identified.





Figure 7. \hat{y} (CC samples) for the external validation set, the dashed horizontal line shows the hard classification threshold of 0.53 (left). Reduced squared residuals Q vs. reduced Hotelling's T^2 plot, the dashed horizontal and vertical lines show the amplitude of the 95% confidence interval for both parameters (right). Samples are marked according to their class (magenta diamonds: cocrystals; cyan circles: binary mixtures).

439 **5.** Conclusion

This study highlighted the ability of a simple PLS–DA model to predict cocrystal formation without any *a priori* knowledge of the specific role of the involved partner molecules. Information deriving from both successful and unsuccessful cocrystallization experiments was used. The major advantage of the proposed methodology relies on the reduction of the experimental effort required for both the synthesis and characterization of new crystalline structures.

The model allows us to predict cocrystallization propensity with a 74% of the predictions in agreement with the experimental results. Considering that the model was obtained on a training set spanning different molecular characteristics, it can be stated that it is suitable for a fairly general applicability.

Indeed, once in possess of the set of chemical descriptors for the molecules of interest, it is sufficient to calculate the absolute value of their difference and perform a linear combination using the pseudo-regression coefficients to obtain a prediction on cocrystal formation. The precalculated values for the set of descriptors comprising 2193 GRAS molecules are available in the Supplementary Material. By applying the proposed methodology, ten new cocrystals were discovered and an additional compound was obtained by chance.

Another figure of merit of the proposed approach is the possibility of understanding through the inspection of PLS weights how the degree of similarity in terms of molecular features of the two partner molecules is correlated with the possibility of obtaining a cocrystal.

On a closing note, we would like to strongly encourage scientists to report failed attempts at cocrystallization along with the technique used, as access to this information could play a pivotal role in refining predictive models, making them less sensitive to selective reporting bias.

458

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464

466 Authorship Contribution Statement

- 467 Fabio Fornari: Conceptualization (equal), Formal Analysis (lead), Methodology (equal), Validation (equal), Visualization
- 468 (equal), Writing Original Draft Preparation (lead), Writing Review & Editing (equal).
- 469 Fabio Montisci: Investigation (equal), Data Curation (lead), Validation (equal), Visualization (equal), Writing Original
- 470 Draft Preparation (supporting), Writing Review & Editing (equal).
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- 480 Writing Review & Editing (supporting).
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484 Review & Editing (equal).

485

486 **Declaration of Competing Interest**

- 487 The authors declare that they have no known competing financial interests or personal relationships that could have appeared
- 488 to influence the work reported in this paper.

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